## MAT5171 Assignment 5

Assignment 5 is based on material from Sections 32-34.

**Exercise** 1 Let  $(S, \mathcal{G})$  be a measurable space and let  $\mu$  be a measure on it. Recall the definition of measure  $(f\mu)$ , where  $f: S \to \mathbb{R}_+$ . Let  $h: S \to \mathbb{R}_+$ . Show that  $h(f\mu) = (hf)\mu$ . Set  $\nu = f\mu$ . then we need to prove  $h\nu = (hf)\mu$ , thus  $\int_B hd\nu = \int_B (hf)d\mu$  for each set B

Furthermore, assume that  $h \in L^1(S, \mathcal{G}, f\mu)$ . Show that  $(f\mu)(h) = \mu(fh)$ .

Hint: Consider first f and h to be indicators, then linear combinations of indicators, then use the argument that any positive function can be approximated by a linear combination.

Note:  $h(f\mu) = (hf)\mu$  is the equality of two measures; while  $(f\mu)(h) = \mu(fh)$  is the equality of two numbers.

**Solution for Exercise 1** We start with a simple function  $h = \sum_{i=1}^{n} a_i 1_{A_i}$ , where  $a_i \ge 0, i = 1, ..., n$ , and  $A_i, i = 1, ..., n$ , is a disjoint partition of S.

Note that  $\nu = f\mu$  is a measure. Hence,  $h(f\mu) = h\nu$  is also a measure. Take  $B \in \mathcal{G}$  to obtain

$$\begin{split} (h(f\mu))(B) &= (h\nu)(B) = \int_{B} h d\nu = \sum_{i=1}^{n} a_{i} \int_{B} 1_{A_{i}} d\nu \\ &= \sum_{i=1}^{n} a_{i} \int 1_{B \cap A_{i}} d\nu = \sum_{i=1}^{n} a_{i} \nu(B \cap A_{i}) \\ &= \sum_{i=1}^{n} a_{i} \int_{B \cap A_{i}} f d\mu = \sum_{i=1}^{n} a_{i} \int_{B} 1_{A_{i}} f d\mu \\ &= \int_{B} \sum_{i=1}^{n} a_{i} 1_{A_{i}} f d\mu = \int_{B} h f d\mu = ((hf)(\mu))(B) \;. \end{split}$$

This proves that  $h(f\mu) = (hf)(\mu)$  for simple functions h.

If function h is measurable, then it can be approximated by an increasing sequence of simple functions. Monotone convergence theorem yields the result for all measurable functions h.

Now, denote again  $\nu = f\mu$ . Then as in the first line of the above computation

$$(f\mu)(h) = \nu(h) = \int h d\nu = \int_S h d\nu = (h(f\mu))(S)$$

and as in the last line of the above computation

$$\mu(fh) = \int fhd\mu = ((hf)(\mu))(S) \; .$$

Since we proved above that the right hand-sides are equal, the left hand sides are equal to each other as well. Note that we proved change of variables formula.

**Exercise** 2 Let  $(\Omega, \mathcal{F}, P)$  be a probability space and let  $\mathcal{H} \subseteq \mathcal{F}$ . Suppose that  $Y : (\Omega, \mathcal{F}, P) \to (S_1, \mathcal{G}_1)$  and  $X : (\Omega, \mathcal{F}, P) \to (S_2, \mathcal{G}_2)$ , where  $(S_1, \mathcal{Y}_1)$ , j = 1, 2, are measurable spaces. Assume that Y is  $\mathcal{H}$ -measurable and X is independent of  $\mathcal{H}$ . Let  $\varphi$  be measurable on the product space  $S_1 \times S_2$  such that  $\varphi(X, Y)$  is integrable.

Prove that  $E[\varphi(X, Y) \mid \mathcal{H}] = g(Y)$ , where  $g(y) = E[\varphi(X, y)]$ .

Solution for Exercise 2 Some preliminary steps:

• Prove that

$$E[\psi(X) \mid \mathcal{H}] = E[\psi(X)]$$

In order to prove it, verify the identity in Theorem 3, following the same steps as in Exercise 3 below.

• Prove that

$$E[\phi(Y) \mid \mathcal{H}] = \phi(Y)$$

In order to prove it, verify the identity in Theorem 3, following the same steps as in Exercise 3 below.

Let  $A \in \mathcal{H}$  be arbitrary. We want to show that

$$E[g(Y)1_A] = E[\varphi(X,Y)1_A]$$

The change of variables formula (applied twice) yields

$$\begin{split} E[g(Y)1_A] &= \int_A g(Y(\omega))P(d\omega) = \int_{Y(A)} g(y)PY^{-1}(dy) \\ &= \int_{Y(A)} E[\varphi(X,y)]PY^{-1}(dy) = \int_{Y(A)} \left( \int_\Omega \varphi(X,y)P(d\omega) \right) PY^{-1}(dy) \\ &= \int_{Y(A)} \left( \int_{X(\Omega)} \varphi(x,y)PX^{-1}(dx) \right) PY^{-1}(dy) \;. \end{split}$$

Here,  $\mu_X = PX^{-1}$  and  $\mu_Y = PY^{-1}$  are the distributions of X and Y, respectively. The assumptions of the theorem imply that X and Y are independent, hence (together with Fubini theorem)

$$E[g(Y)1_A] = \int_{Y(A)} \left( \int_{X(\Omega)} \varphi(x, y) \mu_X(dx) \right) \mu_Y(dy)$$
$$= \int_{Y(A) \times X(\Omega)} \varphi(x, y) (\mu_Y \times \mu_X)(dx, dy) ,$$

where  $\mu_X \times \mu_Y$  is the product measure.

On the other hand, using again change of variables formula,

$$E[\varphi(X,Y)1_A] = \int_A \varphi(X(\omega),Y(\omega))P(d\omega) = \int_{A'} \varphi(x,y)P(X,Y)^{-1}(dx,dy)$$
$$= \int_{A'} \varphi(x,y)(\mu_X \times \mu_Y)(dx,dy) ,$$

where  $P(X,Y)^{-1}$  is the distribution of (X,Y),  $P(X,Y)^{-1} = \mu_X \times \mu_Y$ .

The only one thing that remains to show is to identify  $A^\prime.$  However, note that

$$(X,Y)^{-1}(X(\Omega) \times Y(A)) = \{\omega \in \Omega : X(\omega) \in X(\Omega), Y(\omega) \in Y(A)\} = \{\omega : \omega \in A\} = A$$

Therefore,  $A' = X(\Omega) \times Y(A)$ . Note: You cannot write

$$\int_A \varphi(X,Y) dP = \int \int \varphi(x,y) \mathbf{1}_A \nu(dy) \mu(dx) \; .$$

On the left hand side  $A \subset \Omega$ , while on the right hand side  $A \in ????$ .

**Exercise** 3 Let  $(\Omega, \mathcal{F}, P)$  be a probability space and let  $\mathcal{H} \subseteq \mathcal{G} \subseteq \mathcal{F}$ . Let  $X \in L^1(\Omega, \mathcal{F}, P)$ , Prove the law of the iterated conditional expectations:

$$E[X \mid \mathcal{H}] = E[E[X \mid \mathcal{G}] \mid \mathcal{H}]] .$$

**Solution for Exercise 3** •  $Y = E[X | \mathcal{G}]$ 

- $Z = E[Y \mid \mathcal{H}]$
- By the definition of the conditional expectation we have

$$E[Y1_H] = E[Z1_H]$$

for each  $H \in \mathcal{H}$ .

• Again, by the definition of the conditional expectation we have

$$E[Y1_G] = E[X1_G]$$

for each  $G \in \mathcal{G}$ .

- Since  $\mathcal{H} \subseteq \mathcal{G}$ , for all such sets H we also have  $E[Y1_H] = E[X1_H]$ .
- Thus, for all sets  $H \in \mathcal{H}$  we have  $E[Z1_H] = E[X1_H]$ . Therefore, Z is a version of  $E[X|\mathcal{H}]$ .

**Exercise** 4 Let  $(\Omega, \mathcal{F}, P)$  be a probability space and let  $X \in L^2(\Omega, \mathcal{F}, P)$  (that is, X has the second moment finite). Consider the following minimization problem. Let  $\mathcal{H} \subseteq \mathcal{F}$ . Find a random variable  $Y \in L^2(\Omega, \mathcal{H}, P)$  which is  $\mathcal{H}$ -measurable such that

$$||X - Y||_2 = \inf\{||X - W||_2 : W \in L^2(\Omega, \mathcal{H}, P)\},$$
(1)

where for any random variable Z,  $||Z||_2 = (E[Z^2])^{1/2}$ .

Prove that  $Y = E[X \mid \mathcal{H}]$  solves the minimization problem.

Hint:

- Show that a random variable Y solves (1) if and only if E[(X Y)Z] = 0 for all  $Z \in L^2(\Omega, \mathcal{H}, P)$ . Look when  $||X W||_2^2 \leq ||X Y||_2^2$  holds for  $W = Y + \alpha Z, \ \alpha \in \mathbb{R}$ .
- Show that the random variable Y that satisfies E[(X Y)Z] = 0 for all  $Z \in L^2(\Omega, \mathcal{H}, P)$  is the conditional expectation. Take  $Z(\omega) = 1_H(\omega)$  for  $H \in \mathcal{H}$ . Then

$$E[(X-Y)1_H] = 0$$

for  $H \in \mathcal{H}$ 

**Solution for Exercise 4** Assume that  $Y \in L^2(\Omega, \mathcal{H}, P)$  solves (1). We will deduce its form. Consider  $W = Y + aZ, Z \in L^2(\Omega, \mathcal{H}, P), a \in \mathbb{R}$ . Then

$$0 \ge ||X - Y - aZ||_2^2 - ||X - Y||_2^2 = a^2 E[Z^2] - 2aE[(X - Y)Z].$$

The equation  $0 = a^2 E[Z^2] - 2a E[(X - Y)Z]$  has two roots:

$$a = 0$$
,  $a = 2E[(X - Y)Z]/E[Z^2]$ .

Hence, the expression on the right hand side is negative for some values of a, which is a contradiction. Hence, we must have E[(X - Y)Z] = 0.

Furthermore, if E[(X - Y)Z] = 0 for all  $Z \in L^2(\Omega, \mathcal{H}, P)$ , then we can take  $Z = 1_H, H \in \mathcal{H}$ . We obtain

$$E[X1_H] = E[Y1_H] ,$$

hence the solution to the minimization problem is a version of the conditional expectation.